



Modelling the effect of hydropeaking-induced stranding mortality on Atlantic salmon population abundance

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Complete List of Authors:	Hedger, Richard; Norwegian Institute for Nature Research Sauterleute, Julian; Sweco Norge AS Sundt-Hansen, Line; Norwegian Institute for Nature Research Forseth, Torbjørn; Norwegian Institute for Nature Research Ugedal, Ola; Norwegian Institute for Nature Research Diserud, Ola; Norwegian Institute for Nature Research Bakken, Tor
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1 Modelling the effect of hydropeaking- 2 induced stranding mortality on Atlantic 3 salmon population abundance

4 Effect of hydropeaking stranding mortality on Atlantic salmon 5 population abundance

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8 Richard D. Hedger: Norwegian Institute for Nature Research, NO-7485, Trondheim, Norway.
9 richard.hedger@nina.no. +47 40 46 65 68.

10
11 Julian Sauterleute: Sweco Norway, NO-7030 Trondheim, Norway; SINTEF, NO-7465
12 Trondheim, Norway. julian.sauterleute@sweco.no.

13
14 Line E. Sundt-Hansen: Norwegian Institute for Nature Research, NO-7485, Trondheim,
15 Norway. line.sundt-hansen@nina.no.

16
17 Torbjørn Forseth: Norwegian Institute for Nature Research, NO-7485, Trondheim, Norway.
18 Torbjorn.Forseth@nina.no.

19
20 Ola Ugedal: Norwegian Institute for Nature Research, NO-7485, Trondheim, Norway.
21 Ola.Ugedal@nina.no.

22
23 Ola H. Diserud: Norwegian Institute for Nature Research, NO-7485, Trondheim, Norway.
24 Ola.Diserud@nina.no.

25
26 Tor H. Bakken: SINTEF, NO-7465 Trondheim, Norway. tor.haakon.bakken@sintef.no.

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28 Key words: hydropeaking; stranding mortality; density-dependent mortality; Atlantic salmon;
29 population dynamics

ABSTRACT

Studies of hydropeaking-induced stranding mortality on fish populations have been confined to analysis of empirical data and/or short-term hydraulic-habitat modelling of individual events, and are thus limited as to how they may be used to infer long-term effects in fish populations. In this study, the effects of stranding mortality on an Atlantic salmon population were simulated using an individual-based Atlantic salmon population model with the objective of determining the sensitivity of population dynamics to stranding. It was found that density-dependent mortality (an alternative source of mortality in juvenile Atlantic salmon) partially compensated for stranding mortality, acting as a negative feedback mechanism that dampened change in population abundance. Stranding caused a perturbation in population dynamics, and effects of individual stranding events persisted in time across the life-stages of the population. Effects on population abundance depended on the time of year when stranding was applied, both because of intra-annual changes in stranding mortality probability and because of intra-annual changes in the ability of density-dependent mortality to compensate for stranding mortality. We concluded that empirical measurements of stranding mortality have limited potential for inference of overall effects on the population, and a more dynamic modelling approach, incorporating system feedback, allows for a better modelling of the impact of stranding. Sensitivity analysis showed that population abundance was highly sensitive to density-dependent mortality, and we suggest that this area should be prioritised for further research when investigating the effects of hydropeaking on rivers.

1 Introduction

Hydropeaking is a form of flow regulation used in some regulated rivers which involves initiating frequent, short-term fluctuations in river discharge, executed with the objective of balancing variation in electricity demand through changes in water extraction for powering hydropower turbines (Niu and Insley, 2013). Hydropeaking may detrimentally affect riverine fish populations (Hauer *et al.*, 2017b). The reduction in channel discharge that occurs during down-ramping (short-term reduction of water flow in the regulated watercourse, when water is extracted for use in a turbine) may rapidly dewater shallower areas of the channel, which may lead to fish becoming stranded in dewatered areas (Saltveit *et al.*, 2001; Tuhtan *et al.*, 2012). The risk of this stranding depends on a variety of abiotic and biotic factors including down-ramping speed, season, water temperature, light conditions, fish life-stage and fish behaviour (see Saltveit *et al.*, 2001; Halleraker *et al.*, 2003; Irvine *et al.*, 2009). Stranding may lead to fish mortality, depending on the interaction between the body sizes of the stranded fish, the river bed substrate size, and the length of the stranding event (see Hvidsten, 1985; Bradford, 1997; Nagrodski *et al.*, 2012). In countries reliant on hydropower, hydropeaking may be a pertinent issue with respect to the conservation of riverine fish populations. For example, a majority of Norwegian rivers with high Atlantic salmon (*Salmo salar* L.) production are regulated for hydropower, and hydropower regulation has been a major factor in the loss of 19 stocks (NASCO, 2009). Given that Atlantic salmon population abundances are at historically low levels (Chaput, 2012), research in this area is warranted.

Different approaches have been used to determine how hydropeaking may affect fish populations. Field experiments, involving monitoring fish abundances within net pens placed *in situ* in the river, may be used to directly observe stranding mortality, and determine its sensitivity to different aspects of the flow regime or river characteristics (Saltveit *et al.*, 2001; Halleraker *et al.*, 2003; Irvine *et al.*, 2009) but small sample sizes and often low statistical power make it difficult to generalise results. A statistical approach, involving analysing data collected in rivers affected by or free from hydropeaking may be used to examine aggregate effects of hydropeaking (Liebig *et al.*, 1999; Schmutz *et al.*, 2015), but there is the potential for confounding factors to influence conclusions, and it may be difficult to determine causative pathways as to the manner in which hydropeaking influences fish populations. Hydrodynamic modelling, sometime combined with fish habitat modelling, allows for predicting stranding risk from individual hydropeaking events, and analysing causative relationships under completely controlled conditions (Tiffan *et al.*, 2002; Tuhtan *et al.*, 2012; Hauer *et al.*, 2014). This approach, however, is typically used to model individual or a small number of hydropeaking events, and is limited in regards to the ability for evaluating long-term effects of hydropeaking on fish population dynamics. This is unfortunate because fish populations have complex dynamics: for instance, the increase in fish mortality associated with a stranding event, may reduce fish density, and therefore reduce subsequent density-dependent mortality (an alternative source of mortality that fish experience; see Einum *et al.*, 2006), partially compensating for the stranding event. To model long-term effects of repeated stranding events, some form of population modelling, which incorporates the feedback mechanisms controlling the population, is required.

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3 97 In this study, we investigate the sensitivity of the population dynamics of Atlantic salmon to
4 98 hydropeaking-induced juvenile stranding mortality using an individual based model (IBM)
5 99 approach.

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7 100 1. We begin with an initial simulation where we examine the effect of hydropeaking-
8 101 induced stranding mortality on long-term Atlantic salmon population dynamics. We
9 102 investigate how stranding mortality affects density-dependent mortality (the other
10 103 source of within-river mortality that is simulated in the model).
11 104 2. We then examine the long-term effects of stranding on Atlantic salmon population
12 105 dynamics to determine how long it takes for the equilibrium population to recover from
13 106 the perturbation induced by stranding.
14 107 3. We then examine the relative sensitivity of the Atlantic salmon population to the
15 108 seasonal pattern of hydropeaking.
16 109 4. Finally, we examine the relative sensitivities of the Atlantic salmon population to
17 110 different sources of mortality – both hydropeaking-induced stranding mortality and
18 111 density-dependent mortality – to identify where future research should be focused when
19 112 trying to predict population effects of hydropeaking.
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23 114 **2 Materials and methods**

24 115 25 116 **2.1 Study area**

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27 118 Simulations on the effect of hydropeaking-induced stranding mortality on Atlantic salmon
28 119 population dynamics were parameterised using data from the River Mandalselva, a large
29 120 regulated river (length = 115 km, catchment area = 1800 km², mean annual discharge = 88
30 121 m³ s⁻¹) in Southern Norway (58°2' N, 6°28' E) (see Ugedal *et al.*, 2006). This river was chosen
31 122 because it is one of the most productive Atlantic salmon rivers in Norway that is subject to
32 123 hydropeaking, and has comparatively detailed data on river morphometric characteristics,
33 124 Atlantic salmon population characteristics and flow regulation regime. Hydropower affects
34 125 salmonid populations within this river through a variety of mechanisms, including stranding,
35 126 sedimentation of spawning habitat and obstruction of migration, and it has been estimated
36 127 that the combined effects of hydropower regulation have reduced the total production of
37 128 smolts (early adults that migrate to sea) by 20 – 40% (Ugedal *et al.*, 2006).
38 129

39 130 The watercourse downstream of the outlet from the most downstream hydropower station
40 131 (Laudal hydropower station), comprising a stretch of 19.8 km in length, was modelled. This
41 132 part of the watercourse has a low gradient (change in elevation = 12 m), and meanders
42 133 through alternating sections of floodplains and incised valleys. The channel is predominantly
43 134 composed of glide and walk mesohabitat classes, but there are some rapids. There is a trend
44 135 in substrate size from stones/gravel in the upper part of the watercourse to fine gravel/sand
45 136 in the lower part, with areas of coarser substrate being situated where the river is incising
46 137 through valleys. The channel cross-section is simple, with few islands present, and limited
47 138 departure from a 'U' to 'V' shaped profile. The ratio between peak and base flow in this
48 139 stretch is typically between 1.5 and 2, but may be as high as ≈ 5 in extreme situations.
49 140 Ramping rates during hydropeaking are typically low (mean ≈ 5 cm h⁻¹, max ≈ 10 -20 cm h⁻¹).
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51 142 **2.2 Individual based model (IBM)**

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144 The Atlantic salmon population was modelled using IB-salmon, a spatially-explicit IBM. This
145 model was developed to simulate Atlantic salmon population abundances across all life-
146 stages over long-term (multi-decadal) periods. It was calibrated using time-series data on
147 population characteristics (including juvenile abundance, smolt production, and size at age)
148 for a salmon-bearing river, the River Nausta in western Norway (Hedger *et al.*, 2013a), and
149 has been used to simulate the effects of climate change (Hedger *et al.*, 2013b), hydropeaking
150 (Sauterleute *et al.*, 2016) and habitat remediation (Bustos *et al.*, 2017) on Atlantic salmon
151 populations for several rivers in Norway, including the upper and lower watercourses of the
152 River Mandalselva. It simulates individual Atlantic salmon life-history processes (ontogeny,
153 fecundity, mortality, migration) using heuristic functions derived from field and laboratory
154 experiments. Processes modelled are shown in Hedger *et al.* (2013a). Life-stages modelled
155 are: (i) *eggs*, deposited in the riverbed substrate; (ii) *fry*, juvenile fish in the week of swim-up
156 from the hatched eggs; (iii) *parr*, juvenile fish from one week after swim-up to the spring
157 several years later when individuals are large enough to smoltify; (iv) *smolts*, early adults
158 that have smoltified (in the process of adapting to salt water so that they may travel to sea);
159 (v) *adults at sea*; and (vi) *returning adults* that have returned to the river to spawn. Only parr
160 and later developmental stages are modelled as individual elements; egg and fry abundance
161 are modelled using a group based approach for the sake of computational efficiency.
162 Freshwater processes (growth, mortality, migration) are modelled with a weekly resolution
163 throughout a watercourse that is longitudinally compartmentalised into 50 m long sections;
164 marine processes are modelled in a spatially-integrated fashion with a yearly resolution. The
165 age group of an individual increases on the first week of the year; for example, age 0+ parr
166 (parr in the year of swim-up) become 1+ parr the first week of the following year.

167
168 Parr juveniles are subject to two forms of mortality within the IBM: (i) *density-dependent*
169 *mortality* resulting from biomass exceeding that which can be supported by the river, and (ii)
170 *hydropeaking-induced stranding mortality*.

171
172 (i) Density-dependent mortality is the mortality that occurs when the total biomass of
173 Atlantic salmon parr within a given section of the river exceeds the biomass that can be
174 supported (the carrying capacity, K) of that section. This mortality is termed *density-*
175 *dependent* because it only occurs when parr densities are high (relative to what can be
176 supported). In the example in Figure 1 (left panel), the carrying capacity of the section (K_{sec})
177 is 500 g. A total parr biomass below this threshold will not result in density-dependent
178 mortality: for example, a total biomass of 400 g at time T will give a total biomass of 400 g at
179 time $T+1$. However, when the carrying capacity of the section is exceeded – as a result of
180 recruitment from fry into parr, parr individual body mass growth, parr immigration, and/or a
181 reduction in wetted (water covered) area – biomass surplus to the parr carrying capacity are
182 removed from the section: for example, a total biomass of 800 g at time T will give a total
183 biomass of 500 g (equivalent to the section carrying capacity) at time $T+1$. Surplus parr are
184 subject to a pre-defined *density-dependent mortality probability*; survivors migrate to a
185 neighbouring section. The default is that all survivors migrate downstream – the probability
186 of a surviving individual migrating downstream equals one – although this may be changed
187 according to simulation. Thus, in the example shown, of the 300 g of surplus parr (the
188 difference between the 800 g at time T and the 500 g carrying capacity), some are killed and
189 some migrate out of the section. Removal of parr surplus to the section carrying capacity is
190 size-independent: that is, parr are selected at random until the carrying capacity is no longer

191 exceeded. Density-dependent mortality is modelled as a weekly process, so depends on the
192 mean weekly wetted area, which in turn controls the total carrying capacity of the section
193 each week.

194

195 (ii) Hydropeaking-induced stranding mortality is the mortality that juvenile fish are
196 susceptible to if they find themselves in a dewatered part of the channel on the down-
197 ramping phase of a hydropeaking event. This is density-independent: that is, the probability
198 of an individual experiencing mortality is independent of the number of conspecifics within
199 the dewatered area. Thus, the abundance within a dewatered area at time T+1 is directly
200 proportional to the number at time T (Figure 1, right panel). Stranding mortality occurs
201 within dewatered areas with every hydropeaking event, so the total probability of an
202 individual experiencing stranding mortality within any given week (M) in the IBM simulation
203 is calculated from the assigned stranding mortality probability (S), the proportion of the
204 section that is dewatered (A), and the number of stranding events within that week (n) (see
205 Sauterleute *et al.*, 2016):

206

$$M = 1 - (1 - (S \times A))^n \quad (1)$$

207

208 The assigned stranding mortality probability within the IBM may be specified to be
209 dependent on the individual's life-stage (fry; age 0+ parr; and age 1-4+ parr), the diel period
210 and the season. The rationale for developing the model in this way and the selection of
211 stranding mortality probability values were based on field and laboratory experiments
212 (Saltveit *et al.*, 2001; Halleraker *et al.*, 2003; Harby *et al.*, 2004). Potential limitations of this
213 approach are discussed in Section 4.2.

214

215 Atlantic salmon smolts migrating to sea experience an initial post-smolt mortality, an annual
216 sea mortality, and an annual probability of returning to the river along with a body size
217 drawn from a pre-set size distribution. Returning adults subsequently deposit eggs, so egg
218 deposition is affected by the history of what has occurred in the river (i.e. juvenile survival).

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220 2.3 Parameterising the model

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222 The watercourse was modelled as a 1-D along-channel profile, separated into sections of 50
223 m in downstream length, with each section spanning the entire channel. With regard to
224 modelling the Atlantic salmon population, relevant attributes of the watercourse were (i) the
225 location of spawning redds, and the parr carrying capacity per unit area (g m^{-2}), (ii) weekly
226 discharge and water temperature, (iii) wetted area as a function of weekly discharge (which
227 together with parr carrying capacity per unit area determined density-dependent mortality);
228 and (iv) the proportional area dewatered in a stranding event (which affected stranding
229 mortality).

230

231 Initial egg deposition across the modelled watercourse was estimated according to the
232 number of redds (nests) in each section, determined from survey data (B. Skår, Uni Miljø),
233 with the assumption that one adult female would deposit eggs in each redd. The parr
234 carrying capacity per unit area was determined directly as a function of substrate size, using
235 habitat survey data (H. M. Berger, Norwegian Institute for Water Research).

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237 An intra-annual pattern of discharge and water temperature was generated using data
238 obtained from a gauge at the outlet of the Laudal hydropower station. This pattern was
239 repeated to produce a 50-year time-series of mean weekly discharge and water
240 temperature. The total amount of habitat available in each section (i.e. how much was
241 submerged) depended upon weekly discharge and the channel profile. Atlantic salmon body
242 mass growth, which affected total biomass and therefore density-dependent mortality, was
243 dependent on weekly water temperature.

244
245 Detailed measurements of the along-stream variation in the cross-sectional channel profile
246 have not yet been made for the entire watercourse of the River Mandalselva, and resource
247 limitations prevented us from surveying profiles throughout the entire modelled
248 watercourse. We therefore used an approach of surveying stretches with representative
249 channel types, modelling discharge-wetted width relationships within these stretches using
250 1-D hydraulic modelling, and then applying modified discharge-wetted width relationships
251 derived from this modelling to the entire modelled watercourse, with modifications based
252 on measurements from aerial photographs. The channel profiles of three representative
253 stretches of the river were surveyed by ground measurement: (1) a stretch of the river with
254 a more 'U' shaped channel (12 cross-sectional profiles measured over a 460 m long stretch);
255 (2) a stretch of the river with an intermediate, slightly more 'V' shaped channel (7 profiles
256 measured over a 325 m long stretch); and (3) a stretch of the river with a more 'V' shaped
257 channel (9 profiles measured over a 450 m long stretch). The relationship between discharge
258 and wetted width for each representative stretch was determined using the 1-D hydraulic
259 model HEC-RAS (Hydrological Engineering Center, US Army Corps of Engineers). This model
260 calculates the 1-D St. Venant equation for open channel flow using a four-point implicit finite
261 difference scheme. The hydraulic model was calibrated using measurements of water level
262 in each profile for a single discharge by varying the roughness. Steady-state simulations at
263 different discharges were then used to establish discharge-wetted width relationships for
264 each representative channel profile. Polynomial regression was then used to produced
265 discharge-wetted width curves from the simulated values (Table 1). Each 50 m long section
266 of the watercourse was then classified into one of the channel profiles from analysis of aerial
267 photographs and bank-side observations, and the maximum wetted width of each section
268 was determined from measurements from aerial photographs (imagery obtained from
269 norgebilder.no). Wetted width for each 50 m section and each week of the simulation was
270 determined from mean weekly discharge, using the appropriate channel type relationship
271 ('U'-shaped, intermediate of 'V'-shaped), scaled by the ratio between the maximum wetted
272 width of the 50 m section in question and the maximum wetted width of the representative
273 stretch in question.

274
275 The proportion of each section that was dewatered during a down-ramping event (minimum
276 discharge at down-ramping = $20 \text{ m}^3 \text{ s}^{-1}$) was determined from linear relationships established
277 between wetted widths measured from aerial photographs taken at different discharges (17
278 $- 27 \text{ m}^3 \text{ s}^{-1}$ and $46 - 51 \text{ m}^3 \text{ s}^{-1}$). This allowed for a proportion of the channel dewatered
279 ranging between 0 and 1 according to section. Hydropeaking events were independent of
280 weekly discharge, and consequently the area dewatered in a hydropeaking events was
281 independent of the weekly wetted area (which was dependent on the weekly discharge).
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3 283 Wetted areas (for each week and section) were then read by the IBM to determine weekly
4 284 density-dependent mortality. The proportion of each section that was dewatered in each
5 285 stranding event and the number of events per week were then used to determine stranding
6 286 mortality (equation 1).
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8
9 288 The objective of this study was to use the selected river as a means by which we could
10 289 analyse the sensitivity of the population to stranding, and as such, the intent was not to
11 290 produce an exact simulation of the population within the river. However, simulated
12 291 abundances and fish sizes were compared with those measured *in situ* to confirm that the
13 292 model was creating a biologically feasible population. Estimates of parr abundance were
14 293 obtained from the Norwegian Environment Agency (NVE), based on electrofishing at 7
15 294 stations in October/ November, yearly from 2002 – 2010, in the modelled watercourse.
16 295 Measurements of average fish length according to age group (0+, 1+, 2+ and 3+) were
17 296 available from the NVE for 1998 – 2001. Estimates of smolt production were obtained from
18 297 studies by the Norwegian Institute for Nature Research (Ugedal *et al.*, 2006).
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21 299 **2.4 Analysing the effect of hydropeaking**

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23 301 In all simulations, the IBM simulated weekly processes for a period of 50 years. The first ten
24 302 years were used as a “burn-in” time to generate a full population age distribution. The
25 303 remaining 40 years were used for analysis of the effect of hydropeaking-induced stranding
26 304 on the population dynamics. The primary metrics used to quantify this effect on the
27 305 population were parr abundance and smolt production.
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30 308 Four sets of simulations were run: (i) simulations investigating the effect of hydropeaking on
31 309 long-term Atlantic salmon population dynamics, focusing on compensatory changes in
32 310 density-dependent mortality in response to mortality from stranding; (ii) simulations
33 311 investigating how long it took for the equilibrium population to recover from the
34 312 perturbation induced by hydropeaking; (iii) simulations investigating the effect of different
35 313 seasonal patterns of hydropeaking; and (iv) sensitivity analysis simulations, allowing the
36 314 relative importance of uncertainty in stranding mortality with respect to density-dependent
37 315 mortality to be determined.
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40 318 **1) Effect of hydropeaking on long-term population dynamics.** Atlantic salmon population
41 319 abundance was simulated with a hydropeaking regime which involved down-ramping once
42 320 per day on the weekdays of Monday-Friday (a total of five events per week), with a down-
43 321 ramped discharge of $20 \text{ m}^3 \text{ s}^{-1}$. Fry stranding mortality probability (S_{fry}) was assigned to be
44 322 0.9; parr stranding mortality (S_{parr}) was assigned to be 0.2 (constant across age groups).
45 323 Values for S_{fry} and S_{parr} were in the mid-range of values presented within the literature
46 324 (Saltveit *et al.*, 2001; Halleraker *et al.*, 2003; Harby *et al.*, 2004). For comparison, a control
47 325 scenario was run with no hydropeaking.
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49 327

50 328 **2) Population recovery from perturbation induced by hydropeaking.** Simulations were run
51 329 with hydropeaking occurring over a set range of years (1, 2, 3, ... 10) to explore how the
52 330 recovery of a population on the cessation of hydropeaking was influenced by the length of
53 331 time over which hydropeaking had been applied. Stranding mortalities and weekly
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329 hydropeaking frequencies were the same as in the initial simulation ($S_{\text{fry}} = 0.9$; $S_{\text{parr}} = 0.2$; five
330 hydropeaking events per week).

331

332 **3) Population sensitivity to seasonal pattern of hydropeaking.** Simulations were run with
333 hydropeaking absent except for at specific-four week periods within the year (week of year
334 1-4, 5-6, ... 48-52) to determine the seasonal effect of hydropeaking on the population. For
335 each four-week period, two sets of simulations were run: (i) hydropeaking with a fixed
336 stranding mortality probability that was independent of season ($S_{\text{fry}} = 0.9$; $S_{\text{parr}} = 0.2$); and (ii)
337 hydropeaking with a seasonally-dependent stranding mortality (Table 2). In the latter case, separate simulations were run with seasonally-dependent
338 mortalities consistent with hydropeaking during daytime and hydropeaking during night-
339 time. Assigned seasonally-dependent stranding mortality probabilities were based on
340 information in the literature (Saltveit *et al.*, 2001; Halleraker *et al.*, 2003; Harby *et al.*, 2004).
341 Running simulations with both seasonally-independent and seasonally-dependent
342 mortalities allowed determination of whether seasonal effects on population dynamics were
343 a function of seasonal dependence in stranding mortality probability (e.g. greater parr
344 stranding mortality probability if a stranding event occurred in winter than in summer), or
345 the temporal proximity of the stranding event to the time of smoltification, when juveniles
346 became adults, migrated to sea and were no-longer subject to stranding.
347

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349 **4) Sensitivity analysis: stranding mortality versus density-dependent mortality.** A linear
350 sensitive analysis approach was used to determine the sensitivity of smolt production with
351 respect to: (i) stranding mortality probability of fry, 0+ parr and 1-4+ parr, and (ii) density-
352 dependent mortality of parr. With respect to density dependent mortality, two parameters
353 were analysed: density-dependent mortality probability (the probability that surplus parr
354 died); and the probability of a surviving individual migrating downstream. The latter
355 parameter affected subsequent parr density-dependent mortality indirectly by determining
356 total parr biomass in sections of the river that the parr migrated too. For each simulation,
357 the value of the parameter under investigation was modified around the baseline value (Table 3). The range of values for stranding mortality was based on the range of values that
358 have been established in the literature, although we tested a range of fry stranding mortality
359 probabilities ranging from 0 to 1, reflecting the relative lack of research for this life-stage.
360 Values for density-dependent mortality probability have not been well established so we
361 examined a range from 0.05 (a very low mortality rate for individuals exceeding the carrying
362 capacity) to 1 (all individuals exceeding the carrying capacity experiencing mortality). Values
363 for migration direction have also not been well established, but there is a tendency for
364 individuals to migrate downstream (Beall *et al.*, 1994; Brunsdon *et al.*, 2017). We therefore
365 used a range from 0.5 (50% migrating downstream) to 1 (100% migrating downstream). A
366 linear sensitivity analysis approach was used, rather than a more advanced approach such as
367 Monte Carlo modelling, so as to minimize the number of simulations required to determine
368 sensitivities. This was necessary because of the long run-time required for each simulation.
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370

371 **3 RESULTS**

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373 Populations characteristics simulated by the IBM were similar to those observed in terms of
374 overall parr abundance and body length, and smolt production. Mean overall simulated parr
375 abundances were similar to those observed using electrofishing, although the simulated age

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3 376 distribution was less positively skewed than that observed (Figure 2, left panel). Simulated
4 377 body lengths were similar to, but slightly larger, than those observed (Figure 2, right panel).
5 378 Simulated mean smolt productions were 2.84 smolts 100 m⁻² (no hydropeaking) and 0.51
6 379 smolts 100 m⁻² (five hydropeaking events per week), within the same order of smolt
7
8 380 production estimated by Ugedal *et al.* (2006) which ranged from 1.8 – 3.7 smolts 100 m⁻².
9 381 Given the overall similarity between simulated and observed populations, it was concluded
10 382 that the IBM outputs were satisfactory for investigating the sensitivity of Atlantic salmon
11 383 population dynamics to stranding mortality.
12 384

13 385 **3.1 Effect of hydropeaking on long-term population dynamics**

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16 387 Hydropeaking (5 events per week) reduced parr abundance and smolt production relative to
17 388 the control group of no hydropeaking (Figure 3). The percentage reduction in parr
18 389 abundance from stranding tended to increase with increasing age group. For example,
19 390 stranding resulted in a ≈38.6% reduction in the abundance of 0+ parr, and a ≈44.2%
20 391 reduction in the abundance of 3+ parr. However, the overall effect of hydropeaking on the
21 392 size distribution of the population was small because, whether hydropeaking was applied or
22 393 not, the parr age structure was so strongly positively skewed (dominated by parr from
23 394 younger age groups). Stranding resulted in a ≈76.1 % reduction in production of age 2+
24 395 smolt, and a ≈78.7% in the production of age 3+ smolt.
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27 397 The intra-annual pattern of parr abundance with the application of hydropeaking was
28 398 broadly similar to that in the absence of hydropeaking (Figure 4). In both cases, parr
29 399 abundances peaked mid-year (week of year ≈ 26) due to recruitment from fry, and then
30 400 decreased onwards from the summer peak due to density-dependent mortality resulting
31 401 from the parr carrying capacity being exceeded. Density-dependent mortality rose to an
32 402 initial peak around week of year 32, largely due to increases in biomass due to somatic
33 403 growth, but decreased to low levels by week of year 44. Density-dependent mortality only
34 404 occurred again with the next phase of recruitment into parr the following year.
35 405 Hydropeaking caused stranding mortality throughout the year, regardless of the density of
36 406 parr, resulting in a decline in parr abundance. This decline in parr abundance meant that
37 407 parr carrying capacity was exceeded to a lesser extent, so density-dependent mortality was
38 408 reduced. Consequently, density-dependent mortality was less in the simulation with
39 409 hydropeaking than in the simulation without hydropeaking. The effect of stranding on parr
40 410 abundance was cumulative over successive years, reducing the summer maximum parr
41 411 abundance by nearly two-thirds after six years of stranding (Figure 5).
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44 413 **3.2 Population recovery from perturbation induced by hydropeaking**

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47 415 The rate of recovery of parr abundance to pre-hydropeaking abundance levels after the
48 416 cessation of hydropeaking was strongly dependent on the number of years for which
49 417 hydropeaking was applied (Figure 6, upper panel). For example, with hydropeaking being
50 418 applied for one year only, parr abundance returned and remained at pre-hydropeaking
51 419 levels 11 years after cessation of hydropeaking. In contrast, the effect of consistent
52 420 hydropeaking over a 10 year period was to cause such a reduction in abundance such that
53 421 the population had still not recovered 20 years after hydropeaking had ceased.
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423 The rate of recovery was not consistent from year to year, particularly if hydropeaking had
424 only been applied for a small number of years. For example, when hydropeaking was applied
425 for one year only, there was a reduction in parr abundance in that year (year 11) due to
426 stranding mortality, but this was followed by a peak in parr abundance in the following year
427 (year 12). This peak was the result of new recruitment of age 0+ parr. Older, and larger,
428 individuals had been killed by stranding in the previous year, so there was less pressure on
429 the new recruits of age 0+ parr, and less density-dependent mortality of this group. This,
430 post-hydropeaking, short-term peak in parr abundance caused a peak in smolt production
431 several years later, followed by a peak in egg deposition several years later (year 17) when
432 surviving adults returned from the sea to spawn (Figure 6, lower panel). In contrast, when
433 hydropeaking was applied for an extended period of time, the perturbation pattern was
434 simpler. That is, hydropeaking over successive years reduced recruitment into age 0+ parr
435 over successive years, so no short-term peak in parr abundance occurred, resulting in a long-
436 term reduction in parr abundance before a gradual recovery.
437

438 **3.3 Population sensitivity to seasonal pattern of hydropeaking**

439
440 Smolt production was dependent on the season when hydropeaking was applied, whether
441 simulating with seasonally-independent (Figure 7, left panel) or seasonally-dependent
442 stranding mortality probabilities (Figure 7, right panel). In both cases, stranding during
443 winter and spring (weeks of year 41-20) resulted in a smaller smolt production than
444 stranding in summer and autumn (weeks of year 21-40). Sensitivity to when hydropeaking
445 was applied was greatest for the simulations with a seasonally-dependent stranding
446 mortality probability for daytime hydropeaking, resulting from the greater seasonal range in
447 stranding mortality probability for this hydropeaking regime (
448 Table 2).
449

450 **3.4 Sensitivity analysis: stranding mortality versus density-dependent mortality**

451
452 Smolt production was more sensitive to stranding mortality probability of older age groups
453 (Figure 8a). Smolt production was relatively insensitive to fry stranding mortality probability,
454 with no strong trend throughout the range of stranding mortalities considered (from no
455 mortality to 100% mortality in dewatered areas) with mean annual smolt productions
456 varying between 0.50 and 0.52 smolts 100 m⁻². Smolt production was slightly more sensitive
457 to the stranding mortality probability of age 0+ parr: altering the stranding mortality
458 probability by -50% to +50% around the baseline value caused a change in smolt production
459 of ≈+5% (0.53 smolts 100 m⁻²) to ≈-5% (0.49 smolts 100 m⁻²). In contrast, smolt production
460 was more sensitive to the stranding mortality probability of parr in the 1-4+ age group: a
461 change in stranding mortality from -50% to +50% around the baseline values resulted in a
462 change in smolt production of ≈+20% (0.64 smolts 100 m⁻²) to ≈-10% (0.47 smolts 100 m⁻²).
463

464 Smolt production was sensitive to the density-dependent mortality probability of parr
465 (Figure 8b). For example, altering this probability by -50% to +50% around the baseline value
466 caused a change in smolt production of ≈+30% to ≈-20%. This relationship was non-linear,
467 with changes in the margins of the range of values tested for this parameter (at very low or
468 very high mortality probabilities) having drastic effects on smolt production. In comparison,
469 smolt production was relatively insensitive to the downstream migration probability of

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3 470 surplus surviving parr (parr forced out of the section by the carrying capacity being
4 471 exceeded, which had not experienced subsequent density-dependent mortality). That is,
5 472 smolt production was less sensitive to the direction of movement (upstream or downstream)
6 473 than the mortality caused by carrying capacity being exceeded.
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9 475 **4 Discussion**

10 476
11 477 Overall, we achieved an adequate simulation of population characteristics. The
12 478 underestimate of smolt production under conditions of hydropeaking (both in the initial
13 479 simulation and during sensitivity analysis) may have been caused by the fact that the
14 480 simulated hydropeaking regime involved a higher frequency of hydropeaking events than
15 481 may have been applied in the river. The simulations were run with a strong hydropeaking
16 482 regime (5 events per week throughout the year) to highlight a clear effect. We did not know
17 483 the exact number of hydropeaking events applied in the river, but it is likely that the number
18 484 of events applied would have been less than that simulated. There was a difference between
19 485 simulated juvenile body size distribution and that observed. This may have been caused by
20 486 using a growth function parameterised from data from another river using only temperature
21 487 as a predictor. While growth is strongly related to temperature, other factors such as density
22 488 of conspecifics (Teichert *et al.*, 2010) and food availability (Arnekleiv *et al.*, 2006) may affect
23 489 growth. These properties will likely have differed between the river there the model was
24 490 calibrated and the River Mandalselva.
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28 492 Use of a mechanistic modelling approach offered the advantage that it was possible to
29 493 analyse the relative effect of different parameters (for instance, fry stranding mortality
30 494 versus parr stranding mortality) on hydropeaking-induced stranding on Atlantic salmon
31 495 population dynamics. Additionally, the mechanistic modelling approach allowed for
32 496 modelling long-term processes, involving system lag and feedback. This could not be
33 497 modelled effectively using a coupled hydrodynamic-habitat model because currently
34 498 available models do not include stock-recruitment processes so cannot simulate long-term
35 499 population effects. However, the value of model results depends upon the presuppositions
36 500 involved in model development. In this section, we therefore discuss what the IBM approach
37 501 provides in terms of understanding population dynamics resulting from hydropeaking,
38 502 before addressing model uncertainties, and in particular, uncertainties in the effect of
39 503 stranding mortality in comparison to density-dependent mortality, which may be used to
40 504 define where further research should be focused.
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43 506 **4.1 Population dynamics revealed by the IBM approach**

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45 508 The use of an IBM allowed investigation of how hydropeaking-induced stranding may affect
46 509 Atlantic salmon population dynamics across all life-stages, something that would have been
47 510 difficult to infer from empirical *in situ* measurements or models of single stranding events
48 511 where information obtained is simply the mortality of juvenile life-stages resulting from the
49 512 single event. Firstly, it was evident that the effect of stranding mortality on parr would
50 513 propagate throughout the older life stages. Fewer parr resulted in fewer smolts, which in
51 514 turn reduced the number of returning adults, reduced egg deposition, and thus reduced
52 515 recruitment from fry into parr in future years. Therefore, the effect of stranding may be
53 516 long-term. If parr spend several years in freshwater before moving to sea and then spend
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517 several years at sea before returning to spawn, the effects of stranding on recruitment will
518 be evident years after the last stranding event, and the population may take some time to
519 recover even after a hydropeaking regime has been terminated. Secondly, the IBM approach
520 enabled the exploration of potential hydropeaking-induced temporal patterns in the
521 population dynamics that might be contrary to intuition. For example, hydropeaking for a
522 small number of years actually caused a subsequent short-term increase in parr abundance.
523 This was because hydropeaking reduced total parr biomass, so that small recruits in the
524 following year experienced lower density-dependent mortality. This potential type of
525 phenomenon could not be inferred from an empirical study, where only an instantaneous
526 reduction in population abundance from stranding mortality would be seen.

527
528 The IBM approach also enabled investigation of hydropeaking-induced stranding mortality
529 within the context of the main source of mortality that juvenile Atlantic salmon typically
530 experience in rivers – density-dependent mortality. Stranding mortality reduced population
531 abundance (and population density), so caused a consequent reduction in density-
532 dependent mortality. Therefore, changes in density-dependent mortality may act as a
533 negative feedback mechanism, dampening the effect of hydropeaking. Additionally, it was
534 shown that the specific period within the year when stranding occurs is also critically
535 important for determining overall population effects. Empirical studies have shown that
536 mortality in single stranding events is seasonally dependent (Halleraker *et al.*, 2003): for
537 instance, hydropeaking in winter may lead to greater stranding mortality than hydropeaking
538 during summer. However, by specifying a constant stranding mortality in the model, it was
539 still found that hydropeaking during winter had a greater effect on the population than
540 hydropeaking during summer. This was because the main period of density-dependent
541 mortality occurred during summer, and hydropeaking during summer could be partially
542 compensated for by a reduction in density-dependent mortality. In winter and spring,
543 population abundance was too low for density-dependent mortality to be occurring, so
544 there was no possibility for a compensatory density-dependent response to stranding
545 mortality. Therefore, stranding mortality during these seasons caused a comparatively large
546 reduction in population abundance. It is only through an IBM approach, in which it is
547 possible to assign specified stranding mortality probabilities, that it is possible to see a
548 seasonal effect that is caused by the time period of stranding in respect to when density-
549 dependent mortality occurs.

551 **4.2 Uncertainties in modelling hydropeaking-induced stranding mortality**

552
553 The act of modelling a complex environmental system involves simplification to only include
554 those phenomena which are expected to be the most important. In the current study,
555 potential uncertainties were introduced by (i) a fairly crude characterisation of hydropeaking
556 events, (ii) a simple model of the fish response to down-ramping, (iii) the omission of other
557 potential effects of hydropeaking, and (iv) the omission of potential biological influences.

558
559 Firstly, the modelling of hydropeaking-induced dewatering only characterised the main
560 features of hydropeaking events. With regard to modelling density-dependent mortality and
561 stranding mortality, the watercourse was modelled as a 1-D stretch, using 50 m long sections
562 with a discharge-wetted area relationship determined at the scale of the section. In reality,
563 the exact area dewatered will be dependent on more detailed characteristics of the channel

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3 564 within the section. Following from the work of Sauterleute *et al.* (2016), our hypothesis was
4 565 that the simulation of population dynamics over longitudinal reaches of 50 m in length
5 566 would not be improved by using a more advanced 2-D or 3-D hydraulic model. While a 1-D
6 567 approach may provide acceptable estimates of dewatering for a river with simple channel
7 568 characteristics that is adequately surveyed (Casas-Mulet *et al.*, 2015), a 2-D approach has the
8 569 potential for better modelling of the hydraulics if high resolution data are available (Vozinaki
9 570 *et al.*, 2017). Also, down-ramping in the model affected the entire watercourse with the
10 571 same intensity, whereas in reality, there is usually a downstream decline in the down-
11 572 ramping magnitude (Hauer *et al.*, 2017a). The focus of this study was on analysing how
12 573 Atlantic salmon populations respond to changes in wetted width induced by hydropeaking,
13 574 and the possible feedback mechanisms involved, rather than using the model as a predictive
14 575 tool specifically for the River Mandalselva. As such, a crude method was used to determine
15 576 wetted width as a function of discharge (dependent upon the limited data available). This
16 577 would not have affected our ability to analyse how the population responded to a change in
17 578 wetted area. That is, provided our model was used to analyse population responses to a
18 579 change in wetted area, rather than as a predictive model for how wetted area changes
19 580 according to discharge, errors resulting from the crude 1-D modelling approach will not have
20 581 been an issue.
21 582

22 583 Secondly, stranding mortality within the IBM was modelled simply as the product of the
23 584 dewatered area during the down-ramping cycle of the hydropeaking event (potential varying
24 585 between 0, no dewatered area, and 1, all dewatered) and the assigned stranding mortality
25 586 probability. However, the extent to which fish are affected by dewatering is not fully known.
26 587 Increased movement in response to hydropeaking has been observed (Puffer *et al.*, 2015),
27 588 and there is potential for fish to move to a deeper part of the channel on down-ramping.
28 589 Inconsistencies as to how fish react to stranding may be one of the reasons that stranding
29 590 effects reported in the literature have not always been consistent. For instance, Bradford *et al.*
30 591 *et al.* (1995) and Halleraker *et al.* (2003) reported more stranding during daytime whereas
31 592 Bradford (1997) reported more stranding during night. There is potential to model this using
32 593 a 2D hydrodynamic model, coupled with a 2D individual-based model. However, a model of
33 594 this type that incorporates stock-recruitment relationships so that long-term population
34 595 effects over multiple generations has not yet been developed.
35 596

36 597 Thirdly, hydropeaking may have a large number of potential direct and indirect effects on
37 598 fish populations (Young *et al.*, 2011) other than just the stranding mortality that was
38 599 modelled within the IBM. Hydropeaking may be energetically costly to fish, with subsequent
39 600 effects on growth (Puffer *et al.*, 2015) and over-winter survival (Scruton *et al.*, 2008).
40 601 Dewatering parts of the channel may alter stream ecology by changing physiochemical
41 602 properties such as dissolved oxygen, ionised ammonia, and turbidity. The increase in
42 603 turbidity from flow fluctuation may impair fish vision, reducing prey capture success, and
43 604 reduce respiratory activity through gill abrasion. Hydropeaking may also affect the benthic
44 605 and insect community (Cereghino & Lavandier, 1998; Cereghino *et al.*, 2002; Bruno *et al.*,
45 606 2016) which may affect food webs. Finally, hydropeaking may also change the river
46 607 morphology, with subsequent effects on the river's ecological characteristics (Tuhtan *et al.*,
47 608 2012).
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610 In addition to uncertainties related to hydropeaking-induced mortality, there are properties
611 of Atlantic salmon population dynamics that are not well known. For example, density-
612 dependent mortality is a complex issue, with the strength of the dependence depending
613 upon the degree of interaction between conspecifics from different age groups of the same
614 species, and potential interactions with different fish species. These interactions have not
615 been established robustly. All parr in the IBM were pooled in terms of determining if the
616 parr biomass exceeded parr carrying capacity, but it is feasible that there are situations
617 where there is less interaction between different age groups, and where the use of separate
618 age-specific density-dependent mortalities might be advisable. Another aspect that was
619 difficult to parameterise in the IBM was density-dependent migration of Atlantic salmon
620 juveniles: distances vary within populations, with juveniles being shown to move both
621 upstream and downstream (Erkinaro *et al.*, 1998). Additionally, multiple stranding events
622 occurred within the weekly time-step of the IBM, with the effect of these being integrated
623 across the time-step. Density-dependent mortality was only applied the week following the
624 integration of multiple stranding events, rather than after each individual stranding event. It
625 is possible, therefore, that some of the fine-scale interaction between these two sources of
626 mortality was missed. Finally, compartmentalising the watercourse into spatially-integrated
627 sections of 50 m in length limited the ability to model fine-scale processes. For instance, a
628 reduction in wetted area will have resulted in an increase in density-dependent mortality
629 within the model, on the assumption that all wetted area was used. In reality, it is possible
630 that not all of the section would have been suitable habitat, and that the relationship
631 between changes in wetted area and density-dependent mortality in reality may be more
632 complex.

634 4.3 Sensitivity of the population to stranding and non-stranding effects

636 When examining model sensitivity via a linear sensitivity analysis approach, parr density-
637 dependent mortality probability (which occurred when the carrying capacity was exceeded),
638 had a large effect on smolt production. Quantifying this parameter is difficult, and in reality
639 the level will depend on a wide range of factors such as life-stage, time of year, flow
640 conditions, and channel characteristics. Within the range of parameter values explored, the
641 effect on smolt production was much greater than that for the ranges of juvenile stranding
642 mortality probabilities established in the literature. We infer, therefore, that (with the
643 caveat that a relatively simple sensitivity analysis approach was used) it would be prudent to
644 achieve a better understanding of the biological uncertainties in density-dependent
645 mortality, so that possible effects of stranding can be interpreted in the context of this
646 uncertainty for a better management of Atlantic salmon populations in rivers experiencing
647 hydropeaking.

649 5 Conclusions

651 Using a modelling approach (an IBM incorporating stock-recruitment so that long-term
652 population trends could be simulated) it was possible to elucidate how hydropeaking-
653 induced stranding mortality may affect Atlantic salmon population abundance across
654 multiple generations. Effects of stranding on population abundance were complex, involving
655 compensatory changes in juvenile density-dependent mortality (the main source of mortality
656 that juvenile fishes experience in rivers). Stranding immediately preceding smoltification

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3 657 caused a strong decline in smolt production, because there was less potential for a reduction
4 658 in juvenile density-dependent mortality to compensate for stranding mortality. Stranding
5 659 perturbed population dynamics long after the stranding. Thus, results from individual
6 660 empirical experiment on stranding mortality may not necessarily be reflected in overall
7 661 population effects. Smolt production was more sensitive to density-dependent mortality
8 662 than stranding mortality for the range of values considered. There is a lack of knowledge on
9 663 the magnitude of density-dependent mortality, so it is suggested that this area should be
10 664 prioritised for future research.
11 665

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Tables

Table 1. Relationship between discharge, Q ($\text{m}^3 \text{s}^{-1}$), and wetted width, W (m), established from a polynomial regression fitted to the hydraulic model output.

Channel type	Discharge ($\text{m}^3 \text{s}^{-1}$) wetted width (m) relationship	Discharge range
'U'-shaped	$W = 11.24Q$	$Q < 5$
	$W = 71.93 - \frac{250.12}{Q} + \frac{1587.52}{Q^2} - \frac{3648.27}{Q^3}$	$5 < Q < 70$
	$W = 67.66 + 0.02Q$	$Q \geq 70$
Intermediate	$W = 4.15Q$	$Q < 15$
	$W = 78.07 - \frac{256.45}{Q} + \frac{2404.09}{Q^2} - \frac{31635.11}{Q^3}$	$15 < Q < 120$
	$W = 73.99 + 0.02Q$	$Q \geq 120$
'V'-shaped	$W = 4.28Q$	$Q < 17$
	$W = 114.45 - \frac{476.70}{Q} - \frac{14539.50}{Q^2} + \frac{180507.97}{Q^3}$	$17 < Q < 155$
	$W = 105.08 + 0.04Q$	$Q \geq 150$

Table 2. Seasonal stranding mortality probabilities applied for hydropeaking during night-time and daytime.

Season	Weeks of year	Night-time			Daytime		
		Fry	0+ parr	1-4+ parr	Fry	0+ parr	1-4+ parr
Spring	9-21	0.9	0.20	0.15	0.9	0.20	0.15
Summer	22-34	0.9	0.10	0.05	0.9	0.10	0.05
Autumn	35-48	0.9	0.10	0.05	0.9	0.10	0.05
Winter	48-8	0.9	0.15	0.10	0.9	0.40	0.40

Table 3. Parameters explored in sensitivity analysis. Parameters are (i) stranding mortality probability of fry (S_{fry}); (ii) stranding mortality probability of age 0+ parr ($S_{0+\text{parr}}$); (iii) stranding mortality probability of age 1-4+ parr ($S_{1-4+\text{parr}}$); (iv) density-dependent mortality probability of parr (D_{parr}); and (v) probability of a surviving individual migrating downstream (M_{parr}).

Parameter being analysed	Baseline value	Range	Value of other parameters				
			S_{fry}	$S_{0+\text{parr}}$	$S_{1-4+\text{parr}}$	D_{parr}	M_{parr}
S_{fry}	0.90	(0.00, 0.10, ... 1.00)	NA	0.2	0.2	0.5	1.0
$S_{0+\text{parr}}$	0.20	(0.10, 0.12, ... 0.30)	0.9	NA	0.2	0.5	1.0
$S_{1-4+\text{parr}}$	0.20	(0.10, 0.12, ... 0.30)	0.9	0.2	NA	0.5	1.0
D_{parr}	0.50	(0.05, 0.1, ... 1.0)	0.9	0.2	0.2	NA	1.0
M_{parr}	1.0	(0.5, 0.55, ... 1.0)	0.9	0.2	0.2	0.5	NA

Figure captions

- 851
852
853 Figure 1. Sources of mortality within the IBM: density-dependent mortality for parr within a
854 section (left panel); and hydropeaking-induced stranding mortality for parr within a part of
855 the section dewatered during down-ramping (right panel). In both panels, the thick line
856 shows the relationship between the population at time T and the population at time T+1, the
857 dashed line shows the relationship for an equivalent population for time T and time T+1, and
858 the solid arrows show how a population level at time T is transferred into a population level
859 at time T+1. In the left panel, the dotted arrow shows the total of biomass of parr that is not
860 recruited within the section, some of which will move to another section (through parr
861 migration) and some of which will be removed via density dependent mortality, and K_{sec}
862 shows the carrying capacity of the section (dependent on the weekly wetted area).
863
- 864 Figure 2. Simulated and observed population characteristics: parr abundance (week of year =
865 52) (left panel); and parr body length (right panel). Observed characteristics were derived
866 from data acquired by the Norwegian Environment Agency. The simulated population was
867 generated under conditions of five hydropeaking events per week.
868
- 869 Figure 3. Parr abundance (week of year = 52) and annual smolt production under conditions
870 of no hydropeaking (left panels) and hydropeaking applied five days per week (right panels).
871
- 872 Figure 4. Weekly parr abundance and total parr mortality under conditions of no
873 hydropeaking (left panels) and hydropeaking applied five days per week (right panels) for
874 year-of-simulation 11, the first year after model “burn-in”.
875
- 876 Figure 5. Weekly parr abundance (upper panel) and total parr mortality (lower panel) under
877 conditions of hydropeaking applied five days per week for the first six years after
878 hydropeaking is applied. The beginning of the initial year of hydropeaking (year 11) is
879 indicated by a vertical dashed line.
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- 881 Figure 6. Effect of number of years of hydropeaking on parr abundance (week of year = 52)
882 (upper panel) and egg production (lower panel). The beginning of the initial year of
883 hydropeaking (year 11) is indicated by a vertical dashed line.
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- 885 Figure 7. Effect of seasonal pattern of hydropeaking on smolt production under conditions of
886 seasonally-independent stranding mortality probability ($S_{fry} = 0.9$; $S_{parr} = 0.2$) (left panel) and
887 seasonally-dependent stranding mortality probability (right panel). Seasonally-dependent
888 stranding mortalities are shown in Table 2.
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- 890 Figure 8. Sensitivity of smolt production to (a) stranding mortality probability of fry, age 0+
891 parr, and age 1-4+ parr, and (b) parr density-dependent mortality probability and parr
892 downstream migration probability. Changes in smolt production are calculated as changes
893 from the baseline value (Table 3). Baseline values are shown by the vertical dashed line.

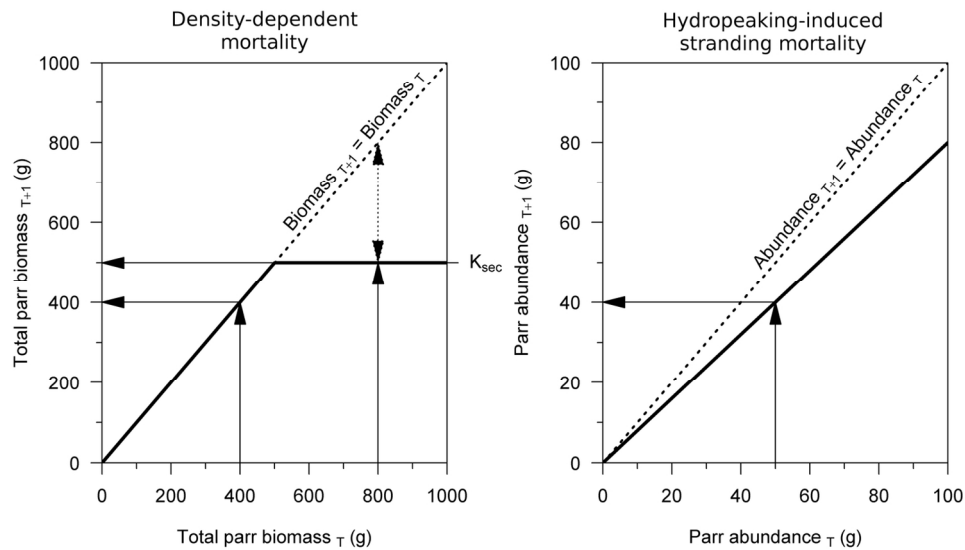


Figure 1. Sources of mortality within the IBM: density-dependent mortality for parr within a section (left panel); and hydropeaking-induced stranding mortality for parr within a part of the section dewatered during down-ramping (right panel). In both panels, the thick line shows the relationship between the population at time T and the population at time $T+1$, the dashed line shows the relationship for an equivalent population for time T and time $T+1$, and the solid arrows show how a population level at time T is transferred into a population level at time $T+1$. In the left panel, the dotted arrow shows the total of biomass of parr that is not recruited within the section, some of which will move to another section (through parr migration) and some of which will be removed via density dependent mortality, and K_{sec} shows the carrying capacity of the section (dependent on the weekly wetted area).

135x80mm (300 x 300 DPI)

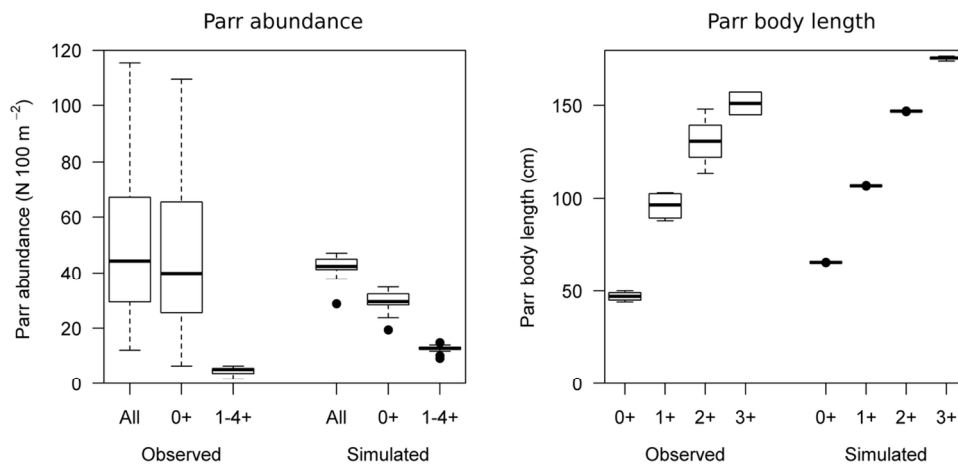


Figure 2. Simulated and observed population characteristics: parr abundance (week of year = 52) (left panel); and parr body length (right panel). Observed characteristics were derived from data acquired by the Norwegian Environment Agency. The simulated population was generated under conditions of five hydropeaking events per week.

114x57mm (300 x 300 DPI)

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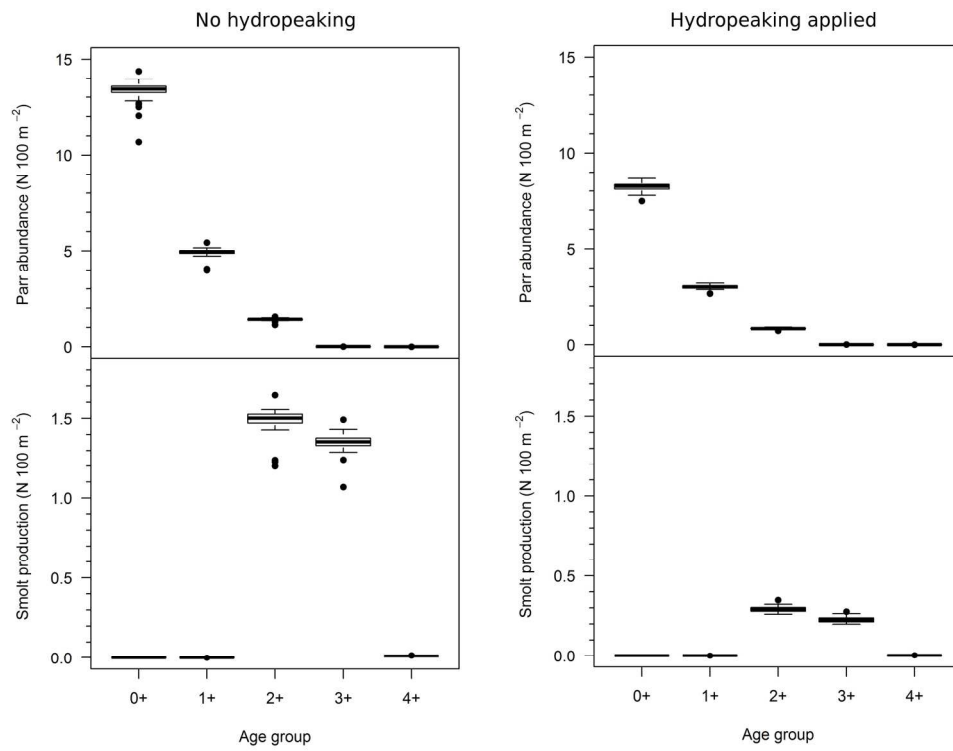


Figure 3. Parr abundance (week of year = 52) and annual smolt production under conditions of no hydropeaking (left panels) and hydropeaking applied five days per week (right panels).

172x133mm (300 x 300 DPI)

view

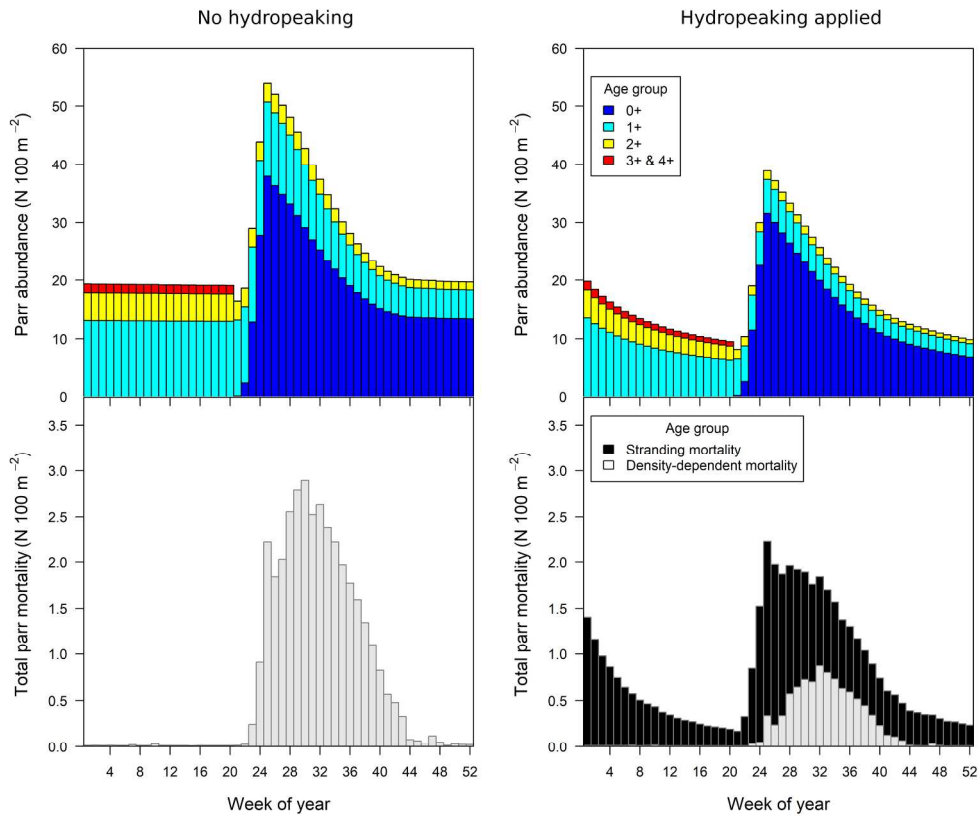


Figure 4. Weekly parr abundance and total parr mortality under conditions of no hydropeaking (left panels) and hydropeaking applied five days per week (right panels) for year-of-simulation 11, the first year after model “burn-in”.

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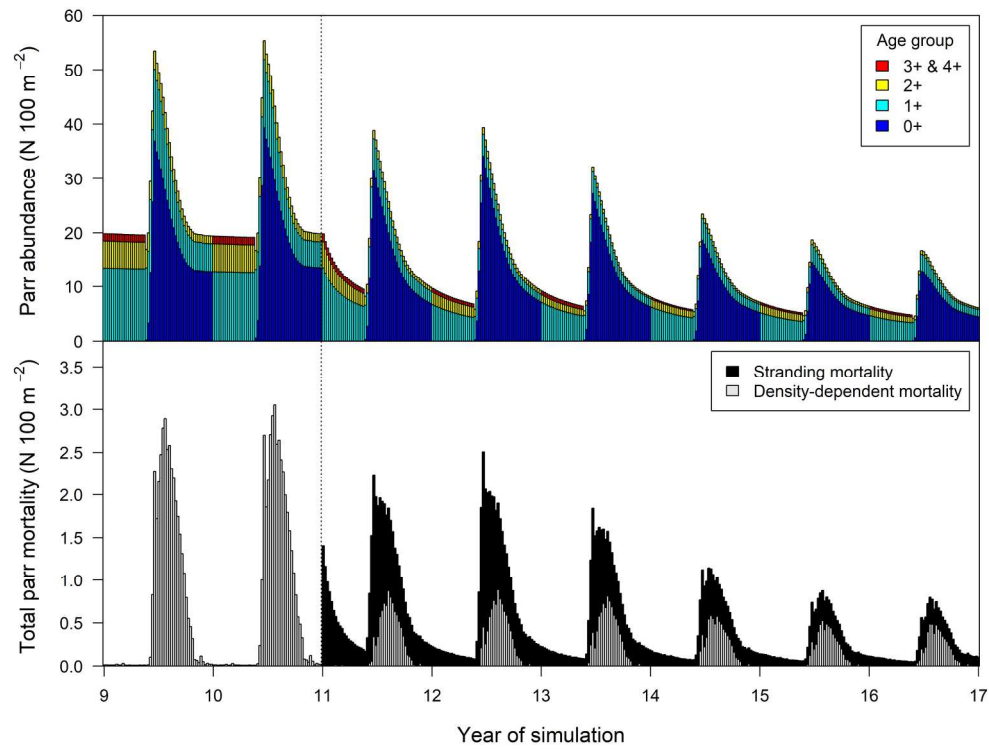


Figure 5. Weekly parr abundance (upper panel) and total parr mortality (lower panel) under conditions of hydropeaking applied five days per week for the first six years after hydropeaking is applied. The beginning of the initial year of hydropeaking (year 11) is indicated by a vertical dashed line.

203x162mm (300 x 300 DPI)

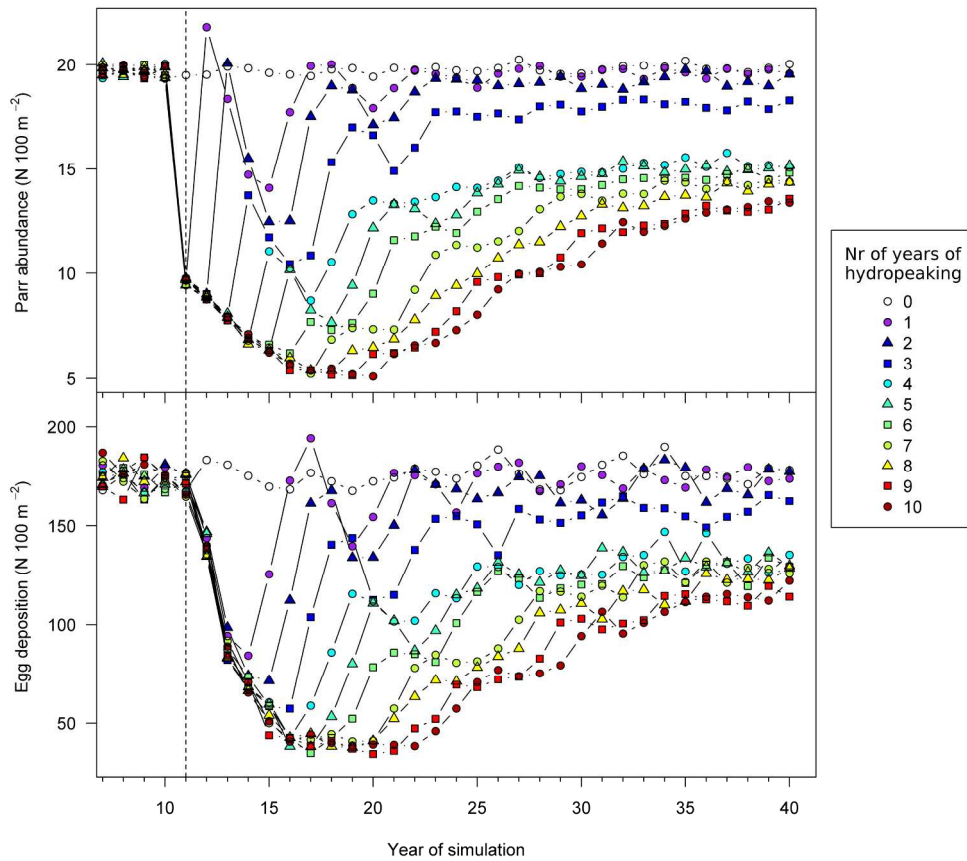


Figure 6. Effect of number of years of hydropeaking on parr abundance (week of year = 52) (upper panel) and egg production (lower panel). The beginning of the initial year of hydropeaking (year 11) is indicated by a vertical dashed line.

227x209mm (300 x 300 DPI)

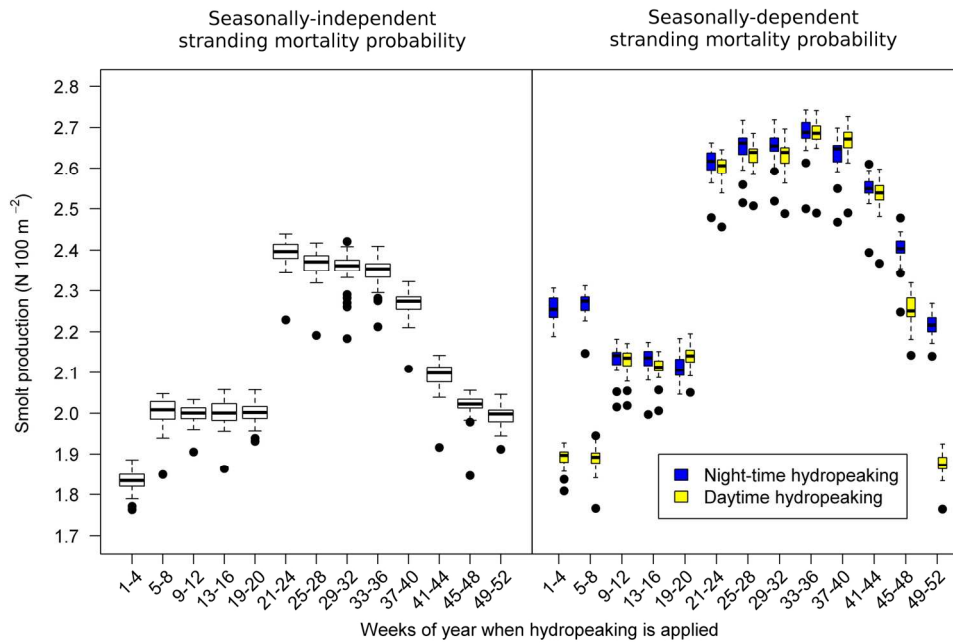


Figure 7. Effect of seasonal pattern of hydropeaking on smolt production under conditions of seasonally-independent stranding mortality probability ($S_{fry} = 0.9$; $S_{parr} = 0.2$) (left panel) and seasonally-dependent stranding mortality probability (right panel). Seasonally-dependent stranding mortalities are shown in Table 2.

160x109mm (300 x 300 DPI)

Review

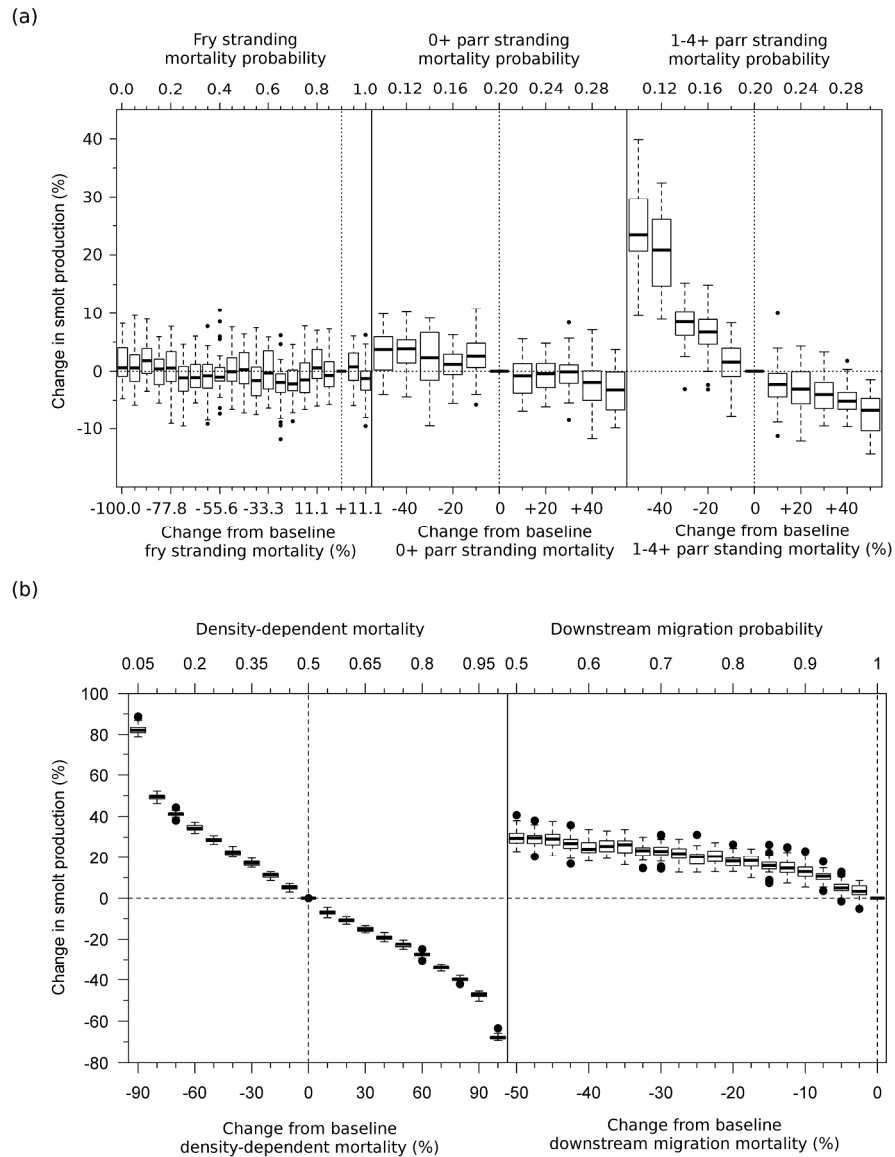


Figure 8. Sensitivity of smolt production to (a) stranding mortality probability of fry, age 0+ parr, and age 1-4+ parr, and (b) parr density-dependent mortality probability and parr downstream migration probability. Changes in smolt production are calculated as changes from the baseline value (Table 3). Baseline values are shown by the vertical dashed line.

302x391mm (300 x 300 DPI)